

RESEARCH ARTICLE

A-optimal design for cubic model without a 3-way effect for mixture experiment

Mahesh Kumar Panda

Department of Statistics, Ravenshaw University, Cuttack 753003, India

Abstract

In this article, we obtain a saturated A-optimal design for the cubic model without a 3-way effect for mixture experiment and get a general formula of the weights. The necessary and sufficient condition of the A-optimality criterion is confirmed by using the corresponding equivalence theorem.

Mathematics Subject Classification (2020). 62K05

Keywords. Saturated design, A-optimal design, cubic model, mixture experiment, equivalence theorem

1. Introduction

A mixture experiment is typically used in formulation and blending experiments. In a mixture experiment, the response is a function of the proportion and not the amount of the ingredients. If a mixture consists of q number of components and proportions of the ingredients are denoted by x_1, x_2, \ldots, x_q , then the experimental region containing these mixture components becomes a (q-1)-dimensional simplex represented as

$$S^{q-1} = \left\{ \mathbf{x} = (x_1, x_2, \dots, x_q)^T \in R^q | \sum_{i=1}^q x_i = 1, 0 \le x_i \le 1, i = 1, 2, \dots, q \right\}.$$
 (1.1)

The construction of an optimal design aims to improve the statistical inference about a particular quantity of interest by appropriately selecting the value of the predictors. These values are so chosen as to minimize the variability of the estimators of the unknown parameters based on a specific criterion. In the literature, different optimal designs are defined like D-, A-, E-, R-, and others. Research work on optimal designs for mixture experiment continued to be an interest to many researchers across the globe. In the last decade, several articles on the construction of optimal designs for mixture models were published [1, 13–17, 19].

The canonical polynomial models are widely used for analyzing mixture data. In particular, Scheffès quadratic model is quite useful in analyzing agricultural and industrial problems [18]. Nevertheless, if the goal is to model the curvature in the interior of a factor space, cubic polynomial models are preferable than quadratic models. Three different

 $Email \ address: \ mahesh 2123 ster@gmail.com$

Received: 12.10.2023; Accepted: 29.09.2024

forms of cubic polynomial models are proposed in the literature i.e., full cubic model, cubic model without a 3-way effect, and special cubic model [2].

In case of D-optimal design, equal weights are allocated to all the support points. In contrast, weights associated with different support points, in general, are not same in case of A-optimal design. Further, the weights assigned to these support points vary with change in the value of q. Therefore, construction of an A-optimal design for higher-degree mixture models involves much more difficulty than the D-optimal design.

Kiefer [7] derived saturated D-optimal designs for full cubic, cubic without a 3-way effect, and special cubic models for a three-component mixture. For other works related to optimal designs for cubic canonical polynomial models for mixture experiment, we can refer to the work of Farrell et al. [3], Mikaeili [11,12], and Lim [10]. Panda and Sahoo [16] obtained a saturated A-optimal design for the full cubic model, cubic model without a three-way effect, and special cubic model for mixture experiment, where q = 3. Recently, Zhu and Hao [20] investigated an A-optimal design for the special cubic mixture model. Notwithstanding such works, the problem of constructing A-optimal design for cubic model without a 3-way effect has not yet been solved. That motivated us to work on the same problem. In this work, we prove that the saturated A-optimal design is supported by the design points of the corresponding D-optimal design. However, the weights assigned to all these design points are not same.

This article is structured as follows: In Section 2, we provide some preliminaries. Section 3 presents a brief discussion on A-optimal design and corresponding equivalence theorem. Section 4 obtains an A-optimal design for the cubic model without a 3-way effect. In Section 5, we obtain the A- efficiency of the corresponding D-optimal design for the discussed model to compare both A-, and D-optimal designs. Finally, we infer the conclusions in Section 6.

2. Preliminaries

In this section, some vectors and matrices are defined that we will use to define the information matrix of the proposed model. These are as follows:

For $q \geq 2$, denote the canonical unit vectors in \mathbb{R}^q by $\mathbf{e}_1, \ldots, \mathbf{e}_q$ and those in $\mathbb{R}^{\binom{q}{2}}$ by \mathbf{E}_{ij} with lexicographically ordered index pairs $(i, j), 1 \leq i < j \leq q$. Here \mathbb{R}^q and $\mathbb{R}^{\binom{q}{2}}$ are the the sets of q- and $\binom{q}{2}$ -dimensional vectors respectively. In the vector \mathbf{e}_i , the i^{th} element is 1 for $i = 1, 2, \ldots, q$, and the remaining elements are 0. Similarly, in the vector \mathbf{E}_{ij} the element at (i, j)th-position is 1 for $1 \leq i < j \leq q$, and the rest of the elements of the vector are 0. For instance, when q = 4:



We denote the identity matrices of dimension q and $\binom{q}{2}$ by \mathbf{U}_1 and \mathbf{W}_1 , respectively. We also denote the vector $\mathbf{1}_q \in R^q$ as $\mathbf{1}_q = (1, \ldots, 1)^{\top}$. Furthermore, we define

$$\mathbf{U_2} = \mathbf{1_q} \mathbf{1_q}^\top - \mathbf{I_q} \in Sym(q),$$

A-optimal design for cubic model without a 3-way effect

$$\begin{split} \mathbf{V_1} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \mathbf{E_{ij}} (\mathbf{e}_i + \mathbf{e}_j)^\top \in R^{\binom{q}{2} \times q}, \\ \mathbf{V_2} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \sum_{\substack{k=1\\k \notin \{i,j\}}}^{q} \mathbf{E_{ij}} \mathbf{e}_k^\top \in R^{\binom{q}{2} \times q}, \\ \mathbf{X_1} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \mathbf{E_{ij}} (\mathbf{e}_j - \mathbf{e}_i)^\top \in R^{\binom{q}{2} \times q}. \end{split}$$

Here Sym(q) and $R^{\binom{q}{2}\times q}$ are the set of symmetric matrices of order q and the set of matrices of order $\binom{q}{2}\times q$, respectively [9].

3. A-optimal design and equivalence theorem

Let us consider a regression model of the form

$$\eta(\mathbf{x}) = \mathbf{f}^{\top}(\mathbf{x})\boldsymbol{\beta}, \quad \mathbf{x} \in S^{q-1}, \tag{3.1}$$

where $\eta(\mathbf{x})$, \mathbf{x} , $\mathbf{f}(\mathbf{x})$ are the expected response, input variable, and regression function respectively. Here we assume that all the responses are independent of each other and have a constant variance. A continuous design $\xi \in \Delta$ have the following form [8]:

$$\xi = \left\{ \begin{matrix} \mathbf{x}_{(1)} & \mathbf{x}_{(2)} & \dots & \mathbf{x}_{(m)} \\ r_1 & r_2 & \dots & r_m \end{matrix} \right\}, \mathbf{x}_{(j)} \in S^{q-1}, 0 < r_j < 1, \sum_{j=1}^m r_j = 1,$$

where $\mathbf{x}_{(1)}, \mathbf{x}_{(2)}, \ldots, \mathbf{x}_{(m)}$ are different design points defined over S^{q-1} ; r_j is the weight associated with the point $x_{(j)}$, j = 1, 2, m. Here Δ is defined as the set of all continuous designs. A non-singular information matrix for a design $\xi \in \Delta$ can be defined as

$$\mathbf{M}(\xi) = \sum_{j=1}^{m} r_j \mathbf{f}(\mathbf{x}_{(j)}) \mathbf{f}^{\top}(\mathbf{x}_{(j)})$$

over S^{q-1} .

Definition 3.1. A design $\xi^* \in \Delta$ with an information matrix $\mathbf{M}(\xi^*)$ for model Equation (3.1) is called A-optimal design if it minimizes $\operatorname{Trace}(\mathbf{M}^{-1}(\xi))$ over the set Δ .

Definition 3.2. A saturated design for any regression model (with p number of parameters) is a design which is supported on exactly p distinct support points [5].

The following equivalence theorem examines the necessary and sufficient conditions of A-optimality over the simplex region S^{q-1} [4].

Theorem 3.3. A continuous design $\xi^* \in \Delta$ is said to be an A-optimal design for the model Equation (3.1) if and only if

$$Max_{\mathbf{x}\in S^{q-1}}d(\mathbf{x},\xi^*) = Trace[\mathbf{M}^{-1}(\xi^*)], \qquad (3.2)$$

where $d(\mathbf{x},\xi) = \mathbf{f}^{\top}(\mathbf{x}) \mathbf{M}^{-2}(\xi) \mathbf{f}(\mathbf{x})$. Furthermore, the supremum exists at the support point of ξ^* .

Selection of support points: The pioneer work on the construction of D-optimal design for the cubic polynomial model without a 3-way effect was due to Kiefer [7]. He considered the initial design $\xi_a \left(0 < a < \frac{1}{2} \right)$ and obtained the D-optimal design for the cubic model without a 3-way effect for the three-component mixture. The design $\xi_a \left(a = \frac{(1-5^{-1/2})}{2} \right)$ that puts equal mass 1/9 to each of the three vertices $\mathbf{x} \longleftrightarrow (1,0,0)$

1761

and six points of the form $\mathbf{x} \longleftrightarrow (a, 1-a, 0)$ is D-optimal design for the cubic model without a 3-way effect. Here $\mathbf{x} \longleftrightarrow (1, 0, 0)$ means the design point (1, 0, 0) and its permutations i.e., (0, 1, 0), (0, 0, 1). Similarly, $\mathbf{x} \longleftrightarrow (a, 1-a, 0)$ means the design point (a, 1-a, 0)and its permutations i.e., (a, 0, 1-a), (0, a, 1-a), (1-a, a, 0), (1-a, 0, a), (0, 1-a, a).

We, therefore, consider the design $\xi_a\left(0 < a < \frac{1}{2}\right)$ that assign weight r_1 to each of the vertices i.e., $\mathbf{x} \longleftrightarrow (1, 0, \dots, 0)$ and weight r_2 to each of the points i.e., $\mathbf{x} \longleftrightarrow (a, 1 - a, \dots, 0)$ to construct a saturated A-optimal design for the cubic model without a 3-way effect.

4. A-optimal design for cubic model without a 3-way effect

In this Section, we obtain A-optimal design for the cubic model without a 3-way effect that can be represented as

$$\eta_1 = \mathbf{f_1}^\top(\mathbf{x})\boldsymbol{\beta_1}$$
$$= \sum_{i=1}^q \beta_i x_i + \sum_{i$$

where $f_1(x)$ and β_1 are column vectors of length q^2 and are defined as

$$\mathbf{f_1}(\mathbf{x}) = (x_1, x_2, \dots, x_q, x_1 x_2, x_1 x_3, \dots, x_{q-1} x_q, x_1 x_2 (x_1 - x_2), \\ x_1 x_3 (x_1 - x_3), \dots, x_{q-1} x_q (x_{q-1} - x_q))^\top, \\ \beta_1 = (\beta_1, \beta_2, \dots, \beta_q, \beta_{12}, \beta_{13}, \dots, \beta_{q-1q}, \delta_{12}, \delta_{13}, \dots, \delta_{q-1q})^\top.$$

The information matrix for the model Equation (4.1) is given by

$$\mathbf{M}(\xi) = \sum_{i=1}^{m} r_i \mathbf{f_1}(\mathbf{x}_{(i)}) \mathbf{f_1}^{\top}(\mathbf{x}_{(i)}).$$
(4.2)

In the next theorem, we obtain the A-optimal allocation ξ_a for the model Equation (4.1).

Theorem 4.1. The design ξ_a that allocates weight $\frac{\sqrt{g_1(a,q)}}{\theta}$ to the support points $\mathbf{x} \longleftrightarrow$ (1,0,...,0) and allocates weight $\frac{\sqrt{g_2(a,q)}}{\theta}$ to the support points $\mathbf{x} \longleftrightarrow$ (a, 1 - a, 0, ..., 0) is the A-optimal allocation, where

$$g_1(a,q) = 1 + \frac{q-1}{2} \frac{1}{a^2(1-a)^2},$$
$$g_2(a,q) = \frac{2a^2 + 1 - 2a}{2a^2(1-a)^2(1-2a)^2},$$

and

$$\theta = q\sqrt{g_1(a,q)} + 2\binom{q}{2}\sqrt{g_2(a,q)}$$

Proof. The information matrix of model Equation (4.1) for the design ξ_a is given by

$$\mathbf{M}(\xi_{a}) = \begin{bmatrix} a_{1}\mathbf{U_{1}} + a_{2}\mathbf{U_{2}} & \frac{a_{2}}{2}\mathbf{V_{1}}^{\top} + 0\mathbf{V_{2}}^{\top} & a_{3}\mathbf{X_{1}}^{\top} \\ \frac{a_{2}}{2}\mathbf{V_{1}} + 0\mathbf{V_{2}} & \frac{a_{2}^{2}}{2r_{2}}\mathbf{W_{1}} & \mathbf{0} \\ a_{3}\mathbf{X_{1}} & \mathbf{0} & a_{4}\mathbf{W_{1}} \end{bmatrix},$$
(4.3)

where **0** is a $\binom{q}{2} \times \binom{q}{2}$ square matrix having each entry equal to 0. Here the coefficients are given by

$$\left. \begin{array}{c} a_1 = r_1 + (q-1)(1+2a(a-1))r_2 \\ a_2 = -2a(a-1)r_2 \\ a_3 = a(a-1)(1-2a)^2r_2 \\ a_4 = 2a^2(1-3a+2a^2)^2r_2 \end{array} \right\}.$$

$$(4.4)$$

The computation of the inverse of the information matrix i.e., $\mathbf{M}^{-1}(\xi_a)$ for the model Equation (4.1) is a quite difficult job. Thus, we compute $Trace(\mathbf{M}^{-1}(\xi_a))$ by considering the sum of the variance of least square estimates of the unknown parameters of the cubic model without a 3-way effect. The least square estimate (LSE) of the parameters of the cubic model without a 3-way effect [2,6] based on design ξ_a are

$$\left. \begin{array}{l} \hat{\beta}_{i} = \hat{\eta}_{i}; \quad i = 1, 2, \dots, q, \\ \hat{\beta}_{ij} = -\frac{1}{2a(1-a)}(\hat{\eta}_{i} + \hat{\eta}_{j}) + \frac{1}{2a(1-a)}(\hat{\eta}_{ij} + \hat{\eta}_{ji}), \quad i < j, \\ \hat{\delta}_{ij} = \frac{1}{2a(1-a)}(\hat{\eta}_{j} - \hat{\eta}_{i}) + \frac{1}{2a(1-a)(1-2a)}(\hat{\eta}_{ji} - \hat{\eta}_{ij}), \quad i < j \end{array} \right\},$$

$$(4.5)$$

where η_i is response to $x_i = 1$, $x_j = 0$, $j \neq i$; $i = 1, 2, \ldots, q$; η_{ij} is response to $x_i = a$, $x_j = 1 - a$, $x_k = 0$, $k \neq i, j$; $i \neq j$.

The variances of the LSEs given in Equation (4.5) are

2

$$\begin{aligned}
& \operatorname{var}(\hat{\beta}_{i}) = \frac{\sigma^{2}}{r_{1}}, \\
& \operatorname{var}(\hat{\beta}_{ij}) = \frac{1}{2a^{2}(1-a)^{2}} \left(\frac{1}{r_{1}} + \frac{1}{r_{2}}\right) \sigma^{2}, \\
& \operatorname{var}(\hat{\delta}_{ij}) = \left(\frac{1}{2a^{2}(1-a)^{2}} \frac{1}{r_{1}} + \frac{1}{2a^{2}(1-a)^{2}(1-2a)^{2}} \frac{1}{r_{2}}\right) \sigma^{2}
\end{aligned}\right\}.$$
(4.6)

Here the real-valued function $Trace(\mathbf{M}^{-1}(\xi_a))$ is proportional to the sum of the variances of LSEs as given in Equation (4.6). Hence, we get

$$T = Trace(\mathbf{M}^{-1}(\xi_a))$$

$$= qvar(\hat{\beta}_i) + \binom{q}{2}var(\hat{\beta}_{ij}) + \binom{q}{2}var(\hat{\delta}_{ij})$$

$$\propto qg_1(a,q)\frac{1}{r_1} + 2\binom{q}{2}g_2(a,q)\frac{1}{r_2},$$
(4.7)

where

$$g_1(a,q) = 1 + \frac{q-1}{2} \frac{1}{a^2(1-a)^2},$$
$$g_2(a,q) = \frac{2a^2 + 1 - 2a}{2a^2(1-a)^2(1-2a)^2}.$$

Now, the problem is to minimize Equation (4.7) subject to the restriction of weights

$$qr_1 + 2\binom{q}{2}r_2 = 1.$$

To solve this problem, we use the Lagrangian multiplier method and set the Lagrangian function as

$$\Psi = Trace(\mathbf{M}^{-1}(\xi_a)) + \lambda \left[qr_1 + 2\binom{q}{2}r_2 - 1 \right].$$

By taking the partial derivatives of Ψ w.r.t r_1 , r_2 , and λ and set them equal to 0, we get

$$-\frac{qg_1(a,q)}{r_1^2} + \lambda q = 0, -2\binom{q}{2}g_2(a,q)\frac{1}{r_1^2} + 2\lambda\binom{q}{2} = 0, qr_1 + 2\binom{q}{2}r_2 - 1 = 0$$

$$(4.8)$$

Solving the system of Equation (4.8), we get

$$r_1 = \frac{\sqrt{g_1(a,q)}}{\theta}, \quad r_2 = \frac{\sqrt{g_2(a,q)}}{\theta}.$$
(4.9)

Hence the theorem is proved.

Next, we find the optimal values of r_1 , r_2 (rounded off to the fourth place of the decimal) and the corresponding value of the $Trace(\mathbf{M}^{-1}(\xi_a))$ for different values of a using Equations (4.7) and (4.9). These values are provided in Table 2 (Appendix I) for different values of q when $3 \leq q \leq 20$. From Table 2, we observe that the design $\xi_a\left(a = \frac{1-5^{-1/2}}{2} = 0.276393\right)$ is a candidate design for the cubic model without a 3-way effect. Let us denote this design by ξ^* .

In the next theorem, we shall prove that the design ξ^* is the A-optimal design for the model Equation (4.1).

Theorem 4.2. The design ξ^* is the A-optimal design for the cubic model without a threeway effect.

Proof. We use appropriate Matlab code by following the steps of Algorithm 1 (Appendix I) and demonstrate numerically that

$$Max_{\mathbf{x}\in S^{q-1}}d(\mathbf{x},\xi^*) = Trace(\mathbf{M}^{-1}(\xi^*)).$$

We also find that equality hold at the support points of the design ξ^* only. The values of $Max_{\mathbf{x}\in S^{q-1}}d(\mathbf{x},\xi^*)$ for different values of q for $3 \leq q \leq 20$ are provided in column (11) of Table 2 (Appendix I). This proves that the design ξ^* is the A-optimal design for the model Equation (4.1) in the class of all designs Δ .

In the next Section, we calculate the A-efficiency of the D-optimal design for different values of q to make a comparison between the A-optimal and D-optimal designs.

5. A-efficiency

Let us denote the A- and D-optimal designs for model Equation (4.1) by ξ_A^* and ξ_D^* respectively. To compare the efficiency, we compute the A- efficiency of the design ξ_D^* for various values of q. Here the A-efficiency of the design ξ_D^* is defined as

$$\Delta_A(\xi_D^*) = \frac{Trace(\mathbf{M}^{-1}(\xi_A^*))}{Trace(\mathbf{M}^{-1}(\xi_D^*))}$$

These efficiency values are displayed in Table 1.

Table 1 shows that the A-efficiency of the D-optimal design continuously decreases as q increases except when q = 4. This indicates that ξ_A^* can significantly improve the ability of parameter estimation.

\overline{q}	$\Delta_A(\xi_D^*)$	
3	99.31%	
4	99.99%	
5	99.58%	
7	98.08%	
10	95.91%	
12	94.70%	
15	93.21%	
20	91.32%	
30	88.84%	
40	87.24%	
50	86.09%	

Table 1. Values of $\Delta_A(\xi_D^*)$ for the cubic model without a 3-way effect for various values of q.

6. Conclusions

The present article obtains a saturated A-optimal design for the cubic model without a 3-way effect with mixture experiment. It is established that the support points of the corresponding D-optimal design are the support points of A-optimal design. Further, the A-optimal design has higher efficiency than the corresponding D-optimal design and this increase in efficiency becomes significant as the number of mixture components increases. For the three-component mixture, we also examine that the derived A-optimal design for the cubic model without a three-way effect are similar to the result obtained by Panda and Sahoo [16].

Finding D- and A-optimal designs for the quartic mixture polynomial model will be interesting. Work in this direction is currently under progess, and we hope to report these findings in a future paper.

Acknowledgements

The author is very grateful to the editor and two anonymous referees for their valuable suggestions.

Author contributions. Not applicable.

Conflict of interest statement. There is no conflict of interest.

Funding. No funding.

Data availability. Not applicable.

References

- M.L. Aggrawal, P. Singh and M.K. Panda, A-optimal designs for an additive cubic model, Stat. Probab. Lett. 81 (2), 259-266, 2011.
- [2] J.A. Cornell, Experiments with Mixtures: Designs, Models, and the Analysis of Mixture Data, John Wiley & Sons, New York, 2002.
- [3] R.H. Farrell, J. Kiefer and J. Walbran, Optimum multivariate designs, Proc. 5th Berkeley Symp. Math. Stat. Probab. 1, 113-139, University of California Press, 1967.

- [4] V.V. Fedorov, Design of experiments for linear optimality criteria, Theory Probab. Appl. 16 (1), 189-195, 1971.
- [5] P. Goos and M. Vandebroek, D-optimal response surface designs in the presence of random block effects, Comput. Stat. Data Anal. 37 (4), 433-453, 2001.
- [6] J.W. Gorman and J. Hinman, E Simplex lattice designs for multicomponent systems, Technometrics 4 (4), 463-487, 1962.
- [7] J. Kiefer, Optimal designs in regression problems II, Ann. Math. Stat. 32 (1), 298-325, 1961.
- [8] J. Kiefer, General equivalence theory for optimum designs (approximate theory), Ann. Statist. 2, 849-879, 1974.
- T. Klein, Invariant symmetric block matrices for the design of mixture experiments, Linear Algebra Appl. 388, 261-278, 2004.
- [10] Y.B. Lim, D-optimal design for cubic polynomial regression on the q-simplex, J. Stat. Plan. Infer. 25 (2), 141-152, 1990.
- [11] F. Mikaeili, D-optimum design for cubic without 3-way effect on the simplex, J. Stat. Plan. Infer. 21 (1), 107-115, 1989.
- [12] F. Mikaeili, D-optimum design for full cubic on q-simplex, J. Stat. Plan. Infer. 35 (1), 121-130, 1993.
- [13] M. Pal and N.K. Mandal, Optimum designs for parameter estimation in mixture experiments with group synergism, Commun. Stat. Theory Methods 50 (9), 2001-2014, 2021.
- [14] M.K. Panda, *R*-optimal designs for canonical polynomial models with mixture experiments, Calcutta Stat. Assoc. Bull. **73** (2), 146-161, 2021.
- [15] M.K. Panda and R.P. Sahoo, *D-optimal designs for Scheffès quadratic mixture model with spline involving two insensitive components*, Int. J. Stat. Reliab. Eng. 9 (1), 108-117, 2022a.
- [16] M.K. Panda and R.P. Sahoo, A-optimal designs for cubic polynomial models with mixture experiments in three components, Stat. Appl. 20 (2), 41-55, 2022b.
- [17] M.K. Panda and R.P. Sahoo, *R-optimal designs for linear log contrast model with mixture experiments*, Commun. Stat. Theory Methods 53 (7), 2355-2368, 2024.
- [18] H. Scheffé, Experiments with mixtures, J. R. Stat. Soc. Ser. B. 20 (2), 344-360, 1958.
- [19] P. Singh and M.K. Panda, Optimal design for second degree K-model for mixture experiments based on weighted simplex centroid design, Metron 69 (3), 251-263, 2011.
- [20] X. Zhu and H. Hao, A-optimal design for the special cubic mixture model, Commun. Stat. Theory Methods 53 (3), 1081-1090, 2024.

APPENDIX I

Algorithm 1. Algorithm to demonstrate Equivalence theorem *Input:*

Step 1. Set the value of q and a = 0.276393. Step 2. Input the values of r_1 from Column (4) and r_2 from Column (7) using Table 2 for the value a = 0.276393. Step 3. Input the column vector

$$\mathbf{f_1}(\mathbf{x}) = (x_1, x_2, \dots, x_q, x_1 x_2, x_1 x_3, \dots, x_{q-1} x_q, x_1 x_2 (x_1 - x_2)),$$

$$x_1x_3(x_1-x_3),\ldots,x_{q-1}x_q(x_{q-1}-x_q))^{+},$$

$$\mathbf{1}_q = (1, \ldots, 1)^{\top}.$$

Step 4. Read the matrices $\mathbf{U_1} = \mathbf{I_q}, \, \mathbf{W_1} = \mathbf{I_{(\underline{q})}}, \, \mathbf{U_2} = \mathbf{1_q} \mathbf{1_q^{\top}} - \mathbf{I_q},$

$$\begin{split} \mathbf{V_1} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \mathbf{E_{ij}}(\mathbf{e}_i + \mathbf{e}_j)^\top, \\ \mathbf{V_2} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \sum_{\substack{k=1\\k \notin \{i,j\}}}^{q} \mathbf{E_{ij}}\mathbf{e}_k^\top, \\ \mathbf{X_1} &= \sum_{\substack{i,j=1\\i < j}}^{\mathbf{q}} \mathbf{E_{ij}}(\mathbf{e}_j - \mathbf{e}_i)^\top. \end{split}$$

Computation:

Step 5. Determine the values of a_1 , a_2 , a_3 , and a_4 by using the following formulae:

$$a_{1} = r_{1} + (q - 1)(1 + 2a(a - 1))r_{2},$$

$$a_{2} = -2a(a - 1)r_{2},$$

$$a_{3} = a(a - 1)(1 - 2a)^{2}r_{2},$$

$$a_{4} = 2a^{2}(1 - 3a + 2a^{2})^{2}r_{2}.$$

Step 6. Obtain the matrix

$$\mathbf{M}(\xi_a) = \begin{bmatrix} a_1 \mathbf{U}_1 + a_2 \mathbf{U}_2 & \frac{a_2}{2} \mathbf{V}_1^\top + 0 \mathbf{V}_2^\top & a_3 \mathbf{X}_1^\top \\ \frac{a_2}{2} \mathbf{V}_1 + 0 \mathbf{V}_2 & \frac{a_2^2}{2r_2} \mathbf{W}_1 & \mathbf{0} \\ a_3 \mathbf{X}_1 & \mathbf{0} & a_4 \mathbf{W}_1 \end{bmatrix}.$$

Step 7. Find the functional form $d(\mathbf{x}, \xi^*)$ where $\mathbf{d}(\mathbf{x}, \xi) = \mathbf{f}^{\top}(\mathbf{x})\mathbf{M}^{-2}(\xi)\mathbf{f}(\mathbf{x})$.

Step 8. Estimate the value of $d(\mathbf{x}, \xi^*)$ at the support points of the design ξ^* . Step 9. Find $Max_{x \in S^{q-1}}d(\mathbf{x}, \xi^*)$.

(11)	$Max_{x\in S^{q-1}}d(\mathbf{x}, \boldsymbol{\xi^*})$					1 2026									1 6 <i>99</i> 0	9009.1								95091 49	<u>04.12002</u>		
(10)	$\operatorname{Trace}[M^{-1}(\xi_a)]$	541680.92	24850.52	7539	3072.69	2708.09	3194.52	5866.44	18037.5	375443.08	1.8517×10^{6}	85285.46	26017.24	10767.32	9663.68	11618	21985.81	69604.85	$1.49018 imes 10^{6}$	4.65444×10^{6}	214974.27	65838.9	$1.20456 imes 10^{6}$	25021.42	30454.69	58730.21	1 20.977
(6)	n	6	6	6	9	6	6	6	6	6	16	16	16	16	16	16	16	16	16	25	25	25	25	25	25	25	л С
(8)	n_2r_2	0.588	0.5988	0.6144	0.6576	0.7062	0.7554	0.8322	0.9072	0.9798	0.636	0.6468	0.6624	0.702	0.7476	0.792	0.8592	0.924	0.984	0.668	0.678	0.692	0.672	0.774	0.814	0.876	0 0 0 0
(2)	r_2	0.098	0.0998	0.1024	0.1096	0.1177	0.1259	0.1387	0.1512	0.1633	0.053	0.0539	0.0552	0.0585	0.0623	0.066	0.0716	0.077	0.082	0.0334	0.0339	0.0346	0.0336	0.0387	0.0407	0.0438	0 0466
(9)	n_2	9	9	9	9	9	9	9	9	9	12	12	12	12	12	12	12	12	12	20	20	20	20	20	20	20	00
(2)	n_1r_1	0.4119	0.4011	0.3855	0.3423	0.2937	0.2445	0.1677	0.0927	0.0201	0.3636	0.3532	0.338	0.2976	0.2524	0.208	0.1404	0.0764	0.0164	0.331	0.3215	0.307	0.3285	0.226	0.185	0.1235	0.067
(4)	r_1	0.1373	0.1337	0.1285	0.1141	0.0979	0.0815	0.0559	0.0309	0.0067	0.0909	0.0883	0.0845	0.0744	0.0631	0.052	0.0351	0.0191	0.0041	0.0662	0.0643	0.0614	0.0657	0.0452	0.037	0.0247	0.0197
(3)	n_1	က	က	က	က	e.	က	က	က	က	4	4	4	4	4	4	4	4	4	ъ	ю	ы	ы	ю	ഹ	ю	ы
(2)	в	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	77
(1)	q					 c	 ဂ									1								 10	ן ה		

1768

				6 6U961	e.e00ee								7 V J U I U I	101204.0								0 10071	0.100411								282340.3
$9.75318 imes 10^{6}$	451409.14	138664.42	58473.41	53603.29	65825.11	128644.14	419749.48	$9.24303 imes10^{6}$	$1.81077 imes 10^7$	839468.86	258483.15	109696.08	101264.61	125204.2	247173.06	813967.75	$1.80756 imes 10^7$	$3.08316 imes10^7$	1.43128×10^{6}	441572.17	188374.89	174881.29	217407.97	432645.07	$1.43509 imes 10^{6}$	$3.20784 imes 10^7$	4.91897×10^{7}	$2.28612 imes 10^6$	706465.16	302698.93	282340.29
36	36	36	36	36	36	36	36	36	49	49	49	49	49	49	49	49	49	64	64	64	64	64	64	64	64	64	81	81	81	81	81
0.693	0.702	0.717	0.753	0.792	0.831	0.888	0.939	0.987	0.7098	0.7224	0.735	0.7686	0.8064	0.8442	0.8988	0.945	0.987	0.728	0.7392	0.7504	0.784	0.8176	0.8568	0.9016	0.952	0.9912	0.7416	0.7488	0.7632	0.792	0.828
0.0231	0.0234	0.0239	0.0251	0.0264	0.0277	0.0296	0.0313	0.0329	0.0169	0.0172	0.0175	0.0183	0.0192	0.0201	0.0214	0.0225	0.0235	0.013	0.0132	0.0134	0.014	0.0146	0.0153	0.0161	0.017	0.0177	0.0103	0.0104	0.0106	0.011	0.0115
30	30	30	30	30	30	30	30	30	42	42	42	42	42	42	42	42	42	56	56	56	56	56	56	56	56	56	72	72	72	72	72
0.3066	0.297	0.2832	0.246	0.207	0.168	0.111	0.06	0.0132	0.2884	0.2786	0.2653	0.2303	0.1925	0.1554	0.1022	0.0546	0.0112	0.272	0.2632	0.2504	0.216	0.18	0.1448	0.096	0.0504	0.0104	0.2592	0.2502	0.2376	0.2052	0.171
0.0511	0.0495	0.0472	0.041	0.0345	0.028	0.0185	0.01	0.0022	0.0412	0.0398	0.0379	0.0329	0.0275	0.0222	0.0146	0.0078	0.0016	0.034	0.0329	0.0313	0.027	0.0225	0.0181	0.012	0.0063	0.0013	0.0288	0.0278	0.0264	0.0228	0.019
9	9	9	9	9	9	9	9	9	2	7	2	7	2	2	2	2	2	∞	∞	x	∞	∞	∞	∞	x	∞	6	6	6	6	6
0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393
				ן ש	0								1 1	_								 0	0								6

	432531.9	635343.5	901654.4
352583.53 706257.02 2.35648×10^{6} 5.29534×10^{7}	$\begin{array}{c} 7.45967 \times 10^7\\ 3.47031 \times 10^6\\ 1.07393 \times 10^6\\ 461870.9\\ 432531.89\\ 542201\\ 1.09206 \times 10^6\\ 3.66164 \times 10^6\\ 8.2644 \times 10^7\\ \end{array}$	$\begin{array}{c} 1.08615 \times 10^{8} \\ 5.05719 \times 10^{6} \\ 1.56694 \times 10^{6} \\ 676099.37 \\ 635343.49 \\ 799047.08 \\ 1.61697 \times 10^{6} \\ 5.44426 \times 10^{6} \\ 1.23335 \times 10^{8} \end{array}$	$\begin{array}{c} 1.52955 \times 10^8 \\ 7.12701 \times 10^6 \\ 2.21067 \times 10^6 \\ 2.21067 \times 10^6 \\ 901654.39 \\ 1.13722 \times 10^6 \\ 2.31071 \times 10^6 \\ 7.80814 \times 10^6 \\ 1.77454 \times 10^8 \end{array}$
81 81 81 81 81	$\begin{array}{c} 100\\ 100\\ 100\\ 100\\ 100\\ 100\\ 100\\ 100$	$\begin{array}{c} 121\\ 121\\ 121\\ 121\\ 121\\ 121\\ 121\\ 121$	$\begin{array}{c} 144\\ 144\\ 144\\ 144\\ 144\\ 144\\ 144\\ 144$
$\begin{array}{c} 0.864 \\ 0.9072 \\ 0.9504 \\ 0.9936 \end{array}$	$\begin{array}{c} 0.756\\ 0.765\\ 0.774\\ 0.801\\ 0.837\\ 0.837\\ 0.873\\ 0.918\\ 0.954\\ 0.99\end{array}$	$\begin{array}{c} 0.759 \\ 0.77 \\ 0.77 \\ 0.781 \\ 0.814 \\ 0.836 \\ 0.836 \\ 0.88 \\ 0.924 \\ 0.957 \\ 1.001 \end{array}$	$\begin{array}{c} 0.7788 \\ 0.7788 \\ 0.792 \\ 0.792 \\ 0.858 \\ 0.8712 \\ 0.8712 \\ 0.924 \\ 0.9636 \\ 0.99 \end{array}$
0.012 0.0126 0.0132 0.0138	$\begin{array}{c} 0.0084\\ 0.0085\\ 0.0086\\ 0.0089\\ 0.0093\\ 0.0097\\ 0.0102\\ 0.0110\\ 0.011\end{array}$	$\begin{array}{c} 0.0069\\ 0.007\\ 0.0074\\ 0.0076\\ 0.008\\ 0.0084\\ 0.0087\\ 0.0091\end{array}$	$\begin{array}{c} 0.0059\\ 0.0059\\ 0.006\\ 0.006\\ 0.0065\\ 0.0066\\ 0.007\\ 0.007\\ 0.0073\\ 0.0075\end{array}$
72 72 72	06 06 06 06 06 06 06 06 06 06 06	$\begin{array}{c} 110\\ 110\\ 110\\ 110\\ 110\\ 110\\ 110\\ 110$	$\begin{array}{c} 132\\ 132\\ 132\\ 132\\ 132\\ 132\\ 132\\ 132\\$
$\begin{array}{c} 0.1368\\ 0.09\\ 0.0486\\ 0.0099\end{array}$	$\begin{array}{c} 0.248\\ 0.239\\ 0.237\\ 0.227\\ 0.196\\ 0.162\\ 0.13\\ 0.03\\ 0.05\\ 0.01\end{array}$	$\begin{array}{c} 0.2387\\ 0.2387\\ 0.2189\\ 0.2189\\ 0.187\\ 0.154\\ 0.154\\ 0.1243\\ 0.0803\\ 0.0429\\ 0.0088\end{array}$	$\begin{array}{c} 0.2292\\ 0.2208\\ 0.21\\ 0.21\\ 0.1476\\ 0.1188\\ 0.0768\\ 0.0408\\ 0.0096\end{array}$
$\begin{array}{c} 0.0152 \\ 0.01 \\ 0.0054 \\ 0.0011 \end{array}$	$\begin{array}{c} 0.0248\\ 0.0239\\ 0.0227\\ 0.0196\\ 0.0162\\ 0.013\\ 0.0085\\ 0.0085\\ 0.0009\end{array}$	$\begin{array}{c} 0.0217\\ 0.0209\\ 0.0199\\ 0.017\\ 0.014\\ 0.0113\\ 0.0113\\ 0.0073\\ 0.0039\\ 0.0008\end{array}$	$\begin{array}{c} 0.0191\\ 0.0184\\ 0.0175\\ 0.0175\\ 0.0175\\ 0.0123\\ 0.0099\\ 0.0009\\ 0.0034\\ 0.0008\end{array}$
0 0 0 0	$\begin{array}{c} 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ 10 \\ 10 $	=======	$\begin{array}{c c} 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 \\ 12 $
$\begin{array}{c} 0.33 \\ 0.4 \\ 0.45 \\ 0.49 \end{array}$	$\begin{array}{c} 0.01\\ 0.05\\ 0.1\\ 0.2\\ \hline 0.2\\ \hline 0.2\\ 0.33\\ 0.4\\ 0.45\\ 0.49\end{array}$	$\begin{array}{c} 0.01 \\ 0.05 \\ 0.1 \\ 0.2 \\ 0.2 \\ \hline 0.33 \\ 0.4 \\ 0.45 \\ 0.49 \end{array}$	$\begin{array}{c} 0.01\\ 0.05\\ 0.1\\ 0.2\\ \hline 0.2\\ \hline 0.2\\ 0.33\\ 0.4\\ 0.45\\ 0.49\end{array}$
	10	11	12

				1.2434 imes	10^{6}								$1.6733 \times$	10^{6}								2.206 imes	10^{6}								2.854 imes
$2.0947 imes 10^{8}$	$5.43339 imes10^7$	$3.03247 imes 10^6$	1.31555×10^6	1.24333×10^{6}	$1.57212 imes 10^{6}$	$3.20588 imes10^{6}$	$1.08672 imes 10^7$	$2.47667 imes10^8$	2.80162×10^{8}	$1.3071 imes 10^7$	$4.06186 imes 10^{6}$	$1.76617 imes 10^6$	$1.67322 imes 10^{6}$	2.12046×10^{6}	$4.33788 imes 10^{6}$	$1.47456 imes 10^7$	3.36885×10^8	$3.67175 imes 10^8$	$1.71399 imes 10^7$	$5.33052 imes10^{6}$	$2.32262 imes10^{6}$	$2.20516 imes 10^{6}$	2.80024×10^{6}	$5.74494 imes 10^{6}$	$1.95774 imes 10^7$	$4.48259 imes 10^{8}$	$4.72795 imes 10^{8}$	$2.20813 imes 10^7$	$6.87227 imes 10^{6}$	$3.00005 imes10^{6}$	$2.85396 imes 10^{6}$
169	169	169	169	169	169	169	169	169	196	196	196	196	196	196	196	196	196	225	225	225	225	225	225	225	225	225	256	256	256	256	256
0.78	0.78	0.7956	0.8268	0.858	0.8892	0.936	0.9516	0.9828	0.7826	0.8008	0.8008	0.8372	0.8736	0.8918	0.9282	0.9646	1.001	0.798	0.798	0.819	0.84	0.882	0.903	0.945	0.966	1.008	0.816	0.816	0.816	0.84	0.888
0.005	0.005	0.0051	0.0053	0.0055	0.0057	0.006	0.0061	0.0063	0.0043	0.0044	0.0044	0.0046	0.0048	0.0049	0.0051	0.0053	0.0055	0.0038	0.0038	0.0039	0.004	0.0042	0.0043	0.0045	0.0046	0.0048	0.0034	0.0034	0.0034	0.0035	0.0037
156	156	156	156	156	156	156	156	156	182	182	182	182	182	182	182	182	182	210	210	210	210	210	210	210	210	210	240	240	240	240	240
0.221	0.2197	0.2041	0.1729	0.143	0.1144	0.0741	0.039	0.0078	0.2156	0.2072	0.1974	0.168	0.1372	0.1106	0.0714	0.0364	0.007	0.2085	0.201	0.1905	0.162	0.1335	0.1065	0.069	0.036	0.0075	0.2032	0.1968	0.1872	0.1584	0.1296
0.017	0.0169	0.0157	0.0133	0.011	0.0088	0.0057	0.003	0.0006	0.0154	0.0148	0.0141	0.012	0.0098	0.0079	0.0051	0.0026	0.0005	0.0139	0.0134	0.0127	0.0108	0.0089	0.0071	0.0046	0.0024	0.0005	0.0127	0.0123	0.0117	0.0099	0.0081
13	13	13	13	13	13	13	13	13	14	14	14	14	14	14	14	14	14	15	15	15	15	15	15	15	15	15	16	16	16	16	16
0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49	0.01	0.05	0.1	0.2	*0.276393
				 									11	14 —								ן די די									16

	3.636×10^{6}	$\begin{array}{c} 4.566 \times \\ 10^6 \end{array}$	5.665×10^{6}
$egin{array}{c} 3.63077 imes 10^6 \ 7.46811 imes 10^6 \ 2.55068 imes 10^7 \ 5.85179 imes 10^8 \end{array}$	$\begin{array}{c} 5.99454 \times 10^8 \\ 2.80096 \times 10^7 \\ 8.72308 \times 10^6 \\ 3.81461 \times 10^6 \\ 3.63539 \times 10^6 \\ 4.63263 \times 10^6 \\ 9.55126 \times 10^6 \\ 3.26884 \times 10^7 \\ 7.5128 \times 10^8 \end{array}$	$\begin{array}{c} 7.49723 \times 10^8 \\ 3.50457 \times 10^7 \\ 1.0921 \times 10^7 \\ 4.78341 \times 10^6 \\ 4.5621 \times 10^6 \\ 5.82772 \times 10^6 \\ 1.20411 \times 10^7 \\ 4.12864 \times 10^7 \\ 9.50436 \times 10^8 \\ 9.50436 \times 10^8 \end{array}$	$\begin{array}{c} 9.26317 \times 10^8 \\ 4.33175 \times 10^7 \\ 1.35064 \times 10^7 \\ 5.92452 \times 10^6 \\ \overline{5.66414 \times 10^6} \\ 7.2392 \times 10^6 \\ 1.4987 \times 10^7 \\ 5.14754 \times 10^7 \\ 5.14754 \times 10^7 \\ 1.18676 \times 10^9 \end{array}$
$256 \\ 256 \\ 256 \\ 256 \\ 256 $	$\begin{array}{c} 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\ 289 \\$	$\begin{array}{r} 324\\ 324\\ 324\\ 324\\ 324\\ 324\\ 324\\ 324\\$	361 361 361 361 361 361 361 361 361
$\begin{array}{c} 0.888\\ 0.936\\ 0.984\\ 1.008\end{array}$	$\begin{array}{c} 0.7888\\ 0.816\\ 0.816\\ 0.8132\\ 0.8432\\ 0.8704\\ 0.8776\\ 0.952\\ 0.952\\ 0.9792\\ 1.0064\end{array}$	$\begin{array}{c} 0.7956\\ 0.8262\\ 0.8262\\ 0.8568\\ 0.8574\\ 0.918\\ 0.918\\ 0.9486\\ 0.9792\\ 1.0098\end{array}$	$\begin{array}{c} 0.8208\\ 0.8208\\ 0.855\\ 0.855\\ 0.8892\\ 0.8892\\ 0.9234\\ 0.9576\\ 0.9576\\ 0.9918\end{array}$
$\begin{array}{c} 0.0037 \\ 0.0039 \\ 0.0041 \\ 0.0042 \end{array}$	$\begin{array}{c} 0.0029\\ 0.003\\ 0.003\\ 0.0031\\ 0.0032\\ 0.0033\\ 0.0035\\ 0.0036\\ 0.0037\end{array}$	$\begin{array}{c} 0.0026\\ 0.0027\\ 0.0027\\ 0.0028\\ 0.0029\\ 0.003\\ 0.0031\\ 0.0033\\ 0.0033\end{array}$	$\begin{array}{c} 0.0024\\ 0.0024\\ 0.0025\\ 0.0026\\ 0.0026\\ 0.0028\\ 0.0028\\ 0.0028\\ 0.0029\end{array}$
$\begin{array}{c} 240 \\ 240 \\ 240 \\ 240 \end{array}$	272 272 272 272 272 272 272 272 272	306 306 306 306 306 306 306 306	$\begin{array}{c} 342\\ 342\\ 342\\ 342\\ 342\\ 342\\ 342\\ 342\\$
$\begin{array}{c} 0.104 \\ 0.0656 \\ 0.0336 \\ 0.0064 \end{array}$	$\begin{array}{c} 0.1989\\ 0.1904\\ 0.1819\\ 0.1547\\ 0.1258\\ 0.1258\\ 0.0646\\ 0.034\\ 0.0068\end{array}$	$\begin{array}{c} 0.1944\\ 0.1872\\ 0.1764\\ 0.1764\\ 0.1494\\ 0.1224\\ 0.0972\\ 0.063\\ 0.0324\\ 0.0054\end{array}$	$\begin{array}{c} 0.1881\\ 0.1805\\ 0.171\\ 0.171\\ 0.1463\\ 0.1197\\ 0.095\\ 0.0608\\ 0.0323\\ 0.0057\end{array}$
$\begin{array}{c} 0.0065\\ 0.0041\\ 0.0021\\ 0.0004\end{array}$	$\begin{array}{c} 0.0117\\ 0.0112\\ 0.0107\\ 0.0091\\ 0.006\\ 0.006\\ 0.0038\\ 0.002\\ 0.0004\\ 0.0004\end{array}$	$\begin{array}{c} 0.0108\\ 0.0104\\ 0.0098\\ 0.0083\\ 0.0068\\ 0.0068\\ 0.0054\\ 0.0035\\ 0.0018\\ 0.0003\end{array}$	$\begin{array}{c} 0.0099\\ 0.0095\\ 0.009\\ 0.0077\\ 0.0063\\ 0.0063\\ 0.005\\ 0.0032\\ 0.0017\\ 0.0003\end{array}$
$16 \\ 16 \\ 16 \\ 16 \\ 16 \\ 16 \\ 16 \\ 16 \\$	$\begin{array}{c} 17\\17\\17\\17\\17\\17\\17\\17\\17\\17\\17\\17\\17\\1$	$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	$ \begin{array}{c} 19\\ 19\\ 19\\ 19\\ 19\\ 19\\ 19\\ 19\\ 19\\ 19\\$
$\begin{array}{c} 0.33 \\ 0.4 \\ 0.45 \\ 0.49 \end{array}$	$\begin{array}{c} 0.01\\ 0.05\\ 0.1\\ 0.2\\ \hline 0.2\\ \hline 0.2\\ 0.33\\ 0.4\\ 0.45\\ 0.49\\ 0.49\end{array}$	$\begin{array}{c} 0.01\\ 0.05\\ 0.1\\ 0.2\\ \hline 0.2\\ \hline 0.2\\ 0.33\\ 0.4\\ 0.45\\ 0.49\end{array}$	$\begin{array}{c} 0.01\\ 0.05\\ 0.1\\ 0.2\\ 0.2\\ \hline 0.2\\ 0.33\\ 0.4\\ 0.45\\ 0.49\end{array}$
	17 —	18	19

				$6.948 \times$	10^{6}			
$1.13209 imes 10^9$	$5.29593 imes10^7$	$1.65214 imes 10^7$	$7.25699 imes 10^{6}$	6.94789×10^{6}	$8.89153 imes 10^{6}$	$1.84415 imes 10^{7}$	$6.344 imes 10^7$	$1.46462 imes 10^{9}$
400	400	400	400	400	400	400	400	400
0.8132	0.836	0.836	0.874	0.874	0.912	0.95	0.988	0.988
0.00214	0.0022	0.0022	0.0023	0.0023	0.0024	0.0025	0.0026	0.0026
380	380	380	380	380	380	380	380	380
0.186	0.178	0.168	0.142	0.116	0.092	0.058	0.03	0.008
0.0093	0.0089	0.0084	0.0071	0.0058	0.0046	0.0029	0.0015	0.0004
20	20	20	20	20		20	20	20
0.01	0.05	0.1	0.2	*0.276393	0.33	0.4	0.45	0.49
				1	70			

* Corresponding D-optimal design $\xi_a \left(a = \frac{1 - 5^{-1/2}}{2} = 0.276393 \dots \right)$

•