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Do energy taxes reduce the carbon footprint? Evidence from Turkey*

Enerji vergileri karbon ayak izini azaltır mı? Türkiye'den kanıtlar

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ÖZ

Bu çalışma, 1994–2017 dönemi Türkiye veri setini kullanarak enerji vergilerinin karbon ayak izi üzerindeki etkisini incelemektedir. Bu çalışmada, değişkenler arasındaki uzun dönemli bir ilişkinin varlığını analiz etmek için Johansen eş-bütünleşme testleri kullanılmaktadır. Ayrıca, uzun dönem eş bütünleşme katsayılarını belirlemek için dinamik en küçük kareler (DOLS) ve tam düzeltilmiş en küçük kareler (FMOLS) tahmin yöntemleri uygulanmaktadır. Johansen eş bütünleşme analizinden elde edilen bulgular, değişkenler arasında uzun dönemli bir eş bütünleşme ilişkisinin olduğunu doğrulamaktadır. Ayrıca, DOLS ve FMOLS sonuçları, enerji vergilerinde meydana gelen bir artışın uzun vadede karbon ayak izini azalttığını ortaya koymaktadır. Bulgular, çevresel vergilerin belirli çevre kalite standartlarına ulaşmak için etkili birer alternatif politika aracı olarak kullanılabilirliğini göstermektedir. Ayrıca, bu tür vergilerin maliyetleri etkilemek suretiyle ekonomik karar birimlerinin çevreye zararlı üretim ve tüketim alışkanlıkları üzerinde etkili olabildiğine dair kanıtlar sunulmaktadır.

ABSTRACT

This study examines the effect of energy taxes on the carbon footprint in Turkey using the data set for 1994–2017. Johansen cointegration tests were used in this study to analyze the existence of a long-term relationship between the variables. Moreover, the estimation methods “dynamic ordinary least squares” (DOLS) and “fully modified ordinary least squares” (FMOLS) were applied to determine the long-term cointegration coefficients. The results of the Johansen cointegration analysis confirm that there is a long-term cointegration relationship between the variables. In addition, the results of DOLS and FMOLS show that an increase in energy taxes reduces the carbon footprint in the long term. In conclusion, the results of the study suggest that environmental taxes can be used as an effective alternative policy tool to achieve certain environmental quality standards. In addition, the results provide evidence that such taxes can influence costs and, by doing so, have an effect on the economic decision units’ habits of environmentally harmful production and consumption.

1. Introduction

With the advent of industrial revolution, energy became one of the important resources for production. In recent years, the effects of globalization movements have also led to an increase in energy use. The structural change that emerged in this process has increased the competitive environment,

and as a result, increasing production and income has become the main goal of economic development (Aydin, Esen and Aydin, 2019). Thus, emerging production and consumption patterns have caused an increase in the demand for energy. As the world population and urbanization increase, human needs and thus environmental damage caused by human-induced pollution also increase (Aydin

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and Esen, 2018; Aydin and Onay, 2020). In this environment, economies focused on increasing their incomes and ignored environmental problems at first. This has caused a number of environmental problems such as the destruction of natural and ecological resources, the more release of carbon emissions in the atmosphere, and the accumulation of non-recyclable waste. Human-induced pollution has reached a level that threatens human health and ecological system.

Environmental degradation has deteriorated gradually due to the increase in economic competition and growth across the world, and this has brought to the forefront the environmental problems and the fight against them (Aydin and Esen, 2017; Aydin, Darici and Kutlu, 2019). Financial tools such as environmental taxes, a large part of which are energy taxes, have become important to combat environmental pollution. Energy tax is a specific taxation and defined as the taxes collected on both transportation and fixed energy products (Gündüz, 2013). The fight against environmental pollution has been the topic of many international conferences, and it has been asserted in these conferences that the problems induced by environmental pollution are caused by carbon and greenhouse gas emissions which are the main cause of climate change. This being the case, the ecosystem has started to be seriously damaged and come under threat due to the overconsumption of resources.

Measures against environmental pollution are also referred to as reform. Green economy is a new production model on a global scale that refers to achieving sustainable development goals while leaving a healthy environment for future generations. Humans leave footprints in the world as a result of their habits of production and consumption throughout life. These footprints are unpredictable in the short term, but may cause permanent damage to nature and therefore to human health in the long term. There are various indicators for sustainable development. One of these indicators is the “ecological footprint”, a new environmental calculation method to measure human demand on nature. Developed in the early 1990s by Rees (1992), Rees and Wackernagel (1994), ecological footprint calculations consist of six main ecological components: carbon footprint, cropland footprint, built-up land footprint, forest land footprint, fishing grounds footprint, and grazing land footprint (WWF, 2012). Carbon footprint accounts for the majority of the ecological footprint. Therefore, ecological footprint also includes greenhouse gas emissions. CO₂ and greenhouse gas emissions, two of the most striking issues of global warming, have affected air pollution and climate change. It has gained importance to reduce emissions in the fight against pollution and climate change. Carbon footprint, expressed in carbon dioxide equivalent, is the measure of the damage caused by human activities to the environment in terms of the amount of greenhouse gas produced.

Since the 1970s the ecological footprint has exceeded and continues to exceed the biocapacity worldwide (WWF,

2012). It is estimated that natural capacity will not be sufficient to meet the demands of humanity in this world order and sustainability will not be possible in such an ecosystem. According to the data of the World Wildlife Fund (WWF), the trends in the use of natural resources in Turkey are similar to the global ones (WWF, 2012). In other words, it is seen that individuals living in Turkey consume natural resources and put pressure on nature at the same rate as the world average. According to the data of WWF (2012), carbon footprint has the largest share (46-49%) in Turkey's total ecological footprint, as it does on a global scale.

The purpose of this study was to examine the effect of energy taxes on the carbon footprint in Turkey between 1994 and 2017. To this end, the existence of a long-term relationship between the variables was analyzed using Johansen cointegration tests. Moreover, the estimation methods “dynamic ordinary least squares” (DOLS) and “fully modified ordinary least squares” (FMOLS) were applied to determine the long-term cointegration coefficients.

To sum up, the rest of the paper is organized as follows. Section 2 reviews the related literature. Section 3 provides the data, the model, and the empirical results. Finally, Section 4 concludes with some discussions and policy suggestions.

2. Literature Review

In many previous studies on the effect of environmental taxes on various pollution indicators (carbon dioxide (CO₂) emissions, sulfur dioxide emissions, exhaust emissions, deforestation, etc.), it has been reported that environmental taxes are important in the development of environmental quality. For example, Manne and Richels (1990), Whalley and Wigle (1991), Alfsen et al. (1995), Nakata and Lamont (2001), Wissema and Dellink (2007), Lu et al. (2010), Miller and Vela (2013), Gemechu et al. (2014), Andersson (2015), Chen et al. (2017), Niu et al. (2018) and Wolde-Rufael and Mulat-Weldemeskel (2021) asserted that environmental taxes had a reducing effect on pollution emissions. Using Fourier Granger causality tests with smooth structural breaks, Aydin (2020) examined the effects of environmental tax revenues on the ecological footprint in 11 OECD countries between 1995 and 2016. He reported that the direction of causality ran from environmental tax revenues to ecological footprint in Germany, Sweden, and Denmark; while in France and Spain, there was a one-way causality running from ecological footprint to environmental tax revenues. Sasmaz (2016) examined a similar subject for EU-15 countries in the period 1995-2012 using panel cointegration tests and Panel FMOLS tests and reported that the increases in environmental tax revenues reduced CO₂ emissions. Similarly, Topal and Günay (2017) examined the data from 53 countries for the period 2000-2014 using static panel data analysis method and reported that environmental taxes had a positive effect on environmental quality. Using the panel smooth transition regression model for the 15 EU

countries, Esen et al. (2021) point out that revenues from environmental taxes significantly reduce ecological deficits after passing a certain threshold level.

On the other hand, results of the studies by Agostini et al. (1992), Bruvoll and Larsen (2004), Gerlagh and Lise (2005), Hotunluoğlu and Tekeli (2007), Lin and Li (2011), Loganathan et al. (2014) and Akkaya and Hepsag (2021) suggest that environmental taxes have a very limited effect in reducing pollutant emissions. Hotunluoğlu and Tekeli (2007) used 18 European countries' data for the period 1995-2003 to examine the effects of carbon taxes on CO₂ emissions. They used the panel least squares method and reported that carbon taxes had a slight effect in reducing emission, that is, they had no significant effect. In their study examining the effects of carbon taxation on CO₂ emissions using time series data for Malaysia from 1974 to 2010, Nanthakumar et al. (2014) asserted that carbon taxation policies were not effective in controlling CO₂ emissions. In a similar study carried out by Akkaya and Hepsag (2021) using Turkey's data for the period 1985-2018, it was reported that taxes on fuel had no effect on CO₂ emissions.

Table 1. Data source and description

Indicator name	Symbol	Unit of measurement	Source
Carbon footprint	carbon	Global hectare (gha) per person	GFN (2021)
Energy taxes revenue	enrtax	% of GDP in million Euros	OECD (2021)
Gross Domestic Product per capita	gdpper	Current US\$	WB (2021)
Energy use	enruse	Oil equivalent per capita (kg)	WB (2021)

The data on carbon footprint was retrieved from the database of Global Footprint Network (GFN, 2021), the data on energy taxes from the database of OECD (Organization for Economic Co-operation and Development) (OECD, 2021),

Table 2. Descriptive statistical analysis for the sample period (1994–2017)

Variables	Observations	Mean	Median	Maximum	Minimum	Std. Dev.
carbon	24	1.616389	1.618395	2.233370	1.066670	0.321150
enrtax	24	2.205625	2.280000	3.197000	0.874000	0.588851
lngdpper	24	8.774088	8.958331	9.442625	7.727684	0.570024
lnenruse	24	7.185880	7.167205	7.504790	6.883960	0.170681

The results of the correlation matrix created to determine the correlation relationship between the variables related to Turkey for the sample period (1994–2017) are presented in Table 3.

Table 3. Result of the correlation matrix test

Variables	carbon	enrtax	lngdpper	lnenruse
carbon	1			
enrtax	0.356261	1		
lngdpper	0.936518	0.500189	1	
lnenruse	0.986301	0.336349	0.926783	1

3. Model, Dataset, and Empirical Results

The effect of energy taxes on carbon footprint was empirically examined in this study. Annual time series data for Turkey from 1994 to 2017 was used to estimate the relationship. The period 1994-2017 was selected based on the availability of data. In the study, the model used to test the relationship between the total tax revenue from energy taxes and the carbon footprint (as a measure of the damage caused by human activities to the environment in terms of the amount of greenhouse gases produced) was defined as follows:

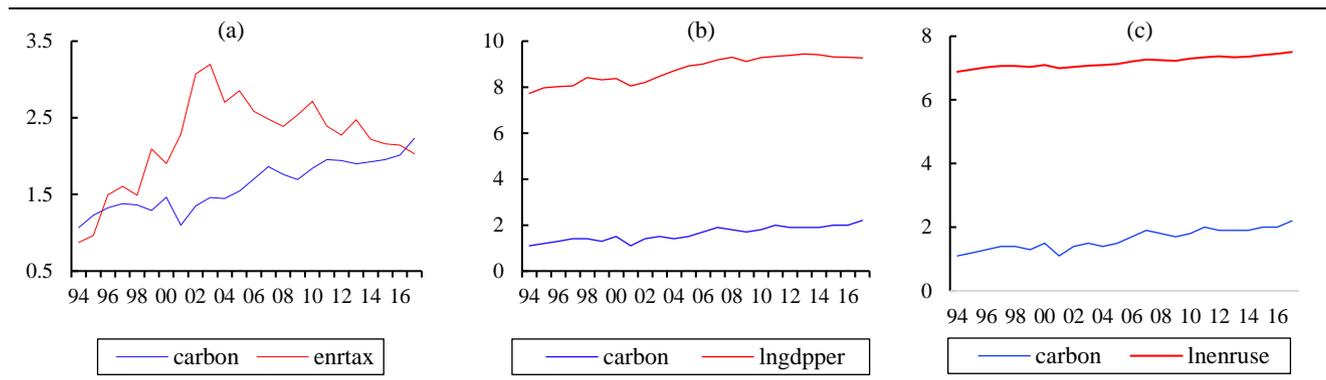
$$carbon_t = \beta_0 + \beta_1 enrtax_t + \beta_2 lngdpper_t + \beta_3 lnenruse_t + \varepsilon_t \quad (1)$$

Where carbon refers to carbon footprint (global hectare (gha) per person), enrtax to the energy taxes revenue (% of GDP in million Euros), lngdpper (natural logarithm of GDP per capita in US\$), lnenruse (natural logarithm of energy use in kg of oil equivalent per capita) and ε to the error term. Table 1 provides a summary of the description and source of the variables.

and the data on per capita income from the database of World Bank (WB). Table 2 shows the descriptive statistics of the data series used in the study.

The correlation coefficients reported in Table 3 indicate the existence of positive and very high correlations between the variables. Besides, the correlation between energy taxes and carbon footprint seems to be quite weak. Figure 1 is important in terms of depicting the high correlation relationships among the variables.

Figure 1. Trends on the relationship between carbon footprint and the variables in Turkey



While graph (b) in Figure 1 shows the parallelism between carbon footprint and per capita income, a similar structure can be clearly seen in graph (c), which reflects the relationship between carbon footprint and per capita energy use. However, in the graph (a), which shows the energy taxes-carbon footprint nexus, it is seen that the relationship between the series changes periodically.

To examine the time series properties of each variable, it must first be determined whether the series are stationary, and if so, at what level they are stationary (Esen, 2012). In time series analysis, in order to obtain empirically

significant relationships between the variables, the series should not contain unit roots, that is, it should be stationary. In general, aggregated time series contain non-stationary behavior (stochastic trend) (Kwiatkowski et al., 1992). Spurious regression problem or invalid statistical inferences may be encountered when studying on non-stationary time series. If this is the case, the results obtained from regression analysis may not reflect the real relationship (Gujarati, 1995; Lin and Brannigan, 2003; Shrestha and Bhatta, 2018). Therefore, in this study, the stationarity of the series was tested using Augmented Dickey-Fuller (ADF) and Phillips-Perron (PP) unit root tests. Table 4 shows the results.

Table 4. Unit root test results

Variables	ADF Unit Root Test				PP Unit Root Test			
	Level		First Difference		Level		First Difference	
	Constant	Constant & Trend	Constant	Constant & Trend	Constant	Constant & Trend	Constant	Constant & Trend
carbon	-0.8069	-3.3026	-6.0127*	-5.8498*	-0.4970	-3.3026	-7.3620*	-7.7942*
enrtax	-2.5048	-1.6503	-4.5159*	-5.3889*	-2.5215	-1.4066	-4.5113*	-6.4444*
lngdpper	-1.7812	-1.3817	-4.5688*	-4.6768*	-1.7832	-1.4708	-4.5688*	-4.6768*
lneruse	-0.4549	-2.4003	-4.7441*	-4.6455*	-0.4153	-2.5249	-4.8960*	-5.0376*

* and ** indicate significance at the 1% and 5% level, respectively.

Based on the results of both ADF and PP unit root tests in Table 4, it was concluded that at $p < 0.05$ significance level, all of the carbon, enrtax, lngdpper and lneruse variables

contained a unit root at level and were stationary when the first differences were considered. Table 5 shows the Dickey-Fuller min-t breakpoint unit root test results.

Table5. Dickey-Fuller min-t breakpoint unit root test results

	Level					
	Intercept		Trend & Intercept		Trend	
	min-t stat.	Breakpoint	min-t stat.	Breakpoint	min-t stat.	Breakpoint
carbon	-4.162	2005	-4.559	2005	-4.031	2002
enrtax	-4.661	2001	-6.663*	2001	-4.301	2003
lngdpper	-3.230	2015	-3.884	2010	-4.053	2013
lneruse	-4.342	2001	-5.126	2005	-4.159	2002
First Difference						
carbon	-6.558*	2001	-6.058*	2001	-6.043*	2017
enrtax	-6.940*	2002	-6.821*	2002	-5.315*	1999
lngdpper	-6.981*	2001	-17.578*	2008	-6.994*	2008
lneruse	-6.567*	2005	-5.569**	2005	-5.152*	1999

The p-values based on Vogelsang (1993) asymptotic one-sided p-values. The lag lengths are determined on the basis of the SIC (Schwarz information criterion). * and ** indicate significance at the 1% and 5% level, respectively.

As in Table 5, when the stationarity status of carbon, enrtax, lngdpper and lneruse series is analyzed by considering

structural breaks, it is concluded that all of the series have unit roots at the level, except for the enrtax variable for all

Models. It is seen that only the *enrtax* series among these variables is stationary in Model 2- trending data with intercept and trend break.

As a result of the unit root tests, it was found that the integration levels of the variables were the same, that is, they were stationary at the same level. Then, it was examined whether there was a long-term relationship between the variables using cointegration tests. Existence of a cointegration between the variables means that there is a long-term relationship. At this stage of the research, Johansen's (1988, 1995) approach was used to examine whether there was a long-term relationship between the variables. The optimal lag length should be determined before performing the Johansen cointegration test. Therefore, the vector autoregressive (VAR) model was estimated to determine the optimal lag length between variables. As a result, the optimal lag length for the Johansen cointegration test was determined as 1 based on the SIC. After determining the optimal lag length, the Johansen cointegration test was used to determine the presence and number of cointegration relationships. Table 6 shows the

Table 6. The results of the cointegration tests and diagnostic tests

H ₀	Johansen cointegration tests				System diagnostics		
	λ_{trace}	%5 c.v.	λ_{max}	%5 c.v.	Serial correlation	Heteroskedasticity	Normality
$r = 0$	69.3475*	54.0790	29.4565*	28.5880	24.71013	177.5726	5.925263
$r \leq 1$	39.8909*	35.1927	24.2103*	22.2996	(0.095)	(0.162)	(0.656)
$r \leq 2$	15.6806	20.2618	9.06671	15.8921			
$r \leq 3$	6.61389	9.16454	6.61389	9.16454			

* denotes statistically rejection of the H₀ in favor of H₁ at the 5% level. The critical values (c.v.) based on MacKinnon-Haug-Michelis (1999). The lag lengths are determined on the basis of the SIC. "r" denotes the number of cointegrating relation. Numbers in () represent p-values.

As for the general model fit, when the reliability of the models was tested using a series of diagnostic tests such as autocorrelation, heteroscedasticity, and normality tests; it was found that the probability values were above 5%. So, there was no autocorrelation or heteroscedasticity problem in the model, and the residuals were normally distributed. As can be seen in Table 6, there is no evidence of serious violations of all diagnostic tests.

After confirming the existence of a long-term relationship between the variables, cointegration parameters should be

Table 7. The results of long-run coefficient estimates based on FMOLS and DOLS for carbon

Variables	DOLS	FMOLS
	Coefficient	Coefficient
<i>enrtax</i>	-0.032212** (0.012511)	-0.024633** (0.011297)
<i>lngdpper</i>	0.241061* (0.036232)	0.080149* (0.026923)
<i>lnenruse</i>	1.038595* (0.115400)	1.622848* (0.083936)
R-squared adj.	0.99	0.97

Standard errors in parentheses. Statistical significance levels: * $p < 0.01$, and ** $p < 0.05$

Table 7 shows the estimation of the long-term relationship between the total tax revenue from energy taxes and the carbon footprint induced by human activities. Based on the DOLS and FMOLS estimation results, it can be asserted that there is a long-term, negative relationship between the

results of the Johansen cointegration test, which is based on Trace and Maximum Eigen-value statistics, for the lag length of 1.

Based on the results in Table 6, it can be asserted that the hypothesis H₀ (there is no cointegration between the variables at the significance level of 5% in both the maximum eigenvalue (λ_{max}) and the trace (λ_{trace}) test statistics) is rejected. So, there are two cointegration relationships between the variables at the 0.05 level in both tests. As can be seen in Table 6, the trace statistic was found to be 69.347, which is above the critical value of 54.079 at the significance level of 5%. Therefore, the null hypothesis ($r = 0$) was rejected at the significance level of 5%. Similarly, also for the maximum eigenvalue test, the null hypothesis of no cointegration was rejected at the significance level of 5% ($\lambda_{max} = 29.456 > 28.588$). These results confirm that there is a cointegration relationship between the variables. In other words, there is a long-term relationship between the energy taxes and the carbon footprint in Turkey.

estimated. The Ordinary Least Squares (OLS) estimator is simple to use for model estimation, but it can cause some problems. For example, in the OLS method, the dynamic effect on the variables that make up the model is not taken into consideration. Moreover, the OLS method may yield biased results in estimating a model with a small sample size. Therefore, in this study, the coefficients of the cointegration vector were analyzed using the DOLS and FMOLS estimation methods. The DOLS and FMOLS estimation results are given in Table 7.

energy taxes and the carbon footprint. In the long term, a one-unit increase in energy taxes decreases the carbon footprint by 0.03 according to the DOLS estimator and by 0.02 according to the FMOLS estimator. The results of both estimators show that energy taxes are an important factor on

the carbon footprint and can contribute to Turkey's environmental development. Moreover, when the DOLS and FMOLS results were examined, it was found that the coefficients of the per capita income and per capita energy use variables were statistically significant at the 0.01 level and positive as expected. Bases on this result, it can be asserted that as the economic activities per capita increase, the carbon footprint induced by these activities also increases in Turkey.

4. Conclusion and Evaluation

Pollution liability, that is, pollution pricing approach is one of the market-based economic approaches that can be actively used to prevent environmental pollution. Pollution liability includes a series of environmental taxes levied to regulate the behavior of producers and consumers. Taxes on energy production and energy products are used as an alternative environmental policy tool in the fight against environmental pollution. The effect of the Turkish energy taxes on carbon footprint was empirically examined in this study. To this end, the relationship was analyzed using the annual data for the period 1994–2017. In the study, cointegration analysis was used to test the existence of a long-term relationship between the variables. In the model, Johansen cointegration tests were applied to test the cointegration relationship, and the DOLS and FMOLS estimators to estimate the long-term coefficients. The empirical results showed that there was a long-term, negative, and statistically significant relationship between the energy taxes and the carbon footprint. It was found that a one-unit increase in the revenues from energy taxes decreased the carbon footprint by 0.03 according to the DOLS estimator and by 0.02 according to the FMOLS estimator. These results show that taxation of energy use provides a price signal because the costs of environmental pollution can be levied on polluters, and this can affect the production and consumption decisions in the context of environmental concerns. In other words, taxes on energy can be used as a policy tool to reduce environmental pollution and/or protect the environment. In conclusion, it can be asserted that such taxes are an appropriate environmental policy tool that can provide incentives for reducing pollutant emissions and thus carbon footprint.

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